Importance sampling simulation of the fork-join queue

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The rare event

 ${S(k) = (S_1(k), S_2(k)) : k = 0, 1, ...}$ is the discrete-time Markov chain analogon of the fork-join queue by embedding at jump times; S(k) represents the backlogs at the queues.



Problem

Estimate by simulation:

$$\gamma_n(x, y, T) = \mathbb{P}(S_1(nT) \ge ny_1 \text{ or } S_2(nT) \ge ny_2 | S(0) = nx),$$

for fixed scaled initial state $x = (x_1, x_2) \in \mathbb{R}^2_+$, fixed scaled threshold $y = (y_1, y_2) \in \mathbb{R}^2_+$, fixed scaled horizon T > 0, and parameter $n \to \infty$.

The fork-join queue

Model

- Poisson (λ) arrivals;
- an arriving job splits in two subjobs;
- two independent single server queues;
- exponential service times with rate μ₁ and μ₂, resp;
- for stability $\lambda < \min(\mu_1, \mu_2)$.



Folklore application: two bathrooms, one for men and one for women, and arrivals of couples.

Original motivation: machine with parallel processors (Hahn&Flatto 1984).

More general: allow individual arrivals (Wright 1992; Shwartz-Weiss

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The rarity set

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The scaled set of interest:

$$D = \{\eta \in \mathbb{R}^2_+ : \eta_1 \ge y_1 \text{ or } \eta_2 \ge y_2\},$$

i.e.

$$\gamma_n(x, y, T) = \mathbb{P}\left(S(nT)/n \in D | S(0) = nx\right).$$

The difficulty for importance sampling is twofold:

- (i). the rarity set is not convex (Dupuis&Wang 2007);
- (ii). the rarity set cannot decomposed in two disjoint sets such that the separate probabilities are estimated by efficient importance sampling estimators (Glassermann&Wang 1997).

A sample path of the fork-join queue

Face-homogeneous random walk

The fork-join queue is a face-homogeneous random walk on \mathbb{Z}^2_+ with four faces.



The transition probabilities $p_{s,s+d}$ are constant for *s* in the same face F_{Λ} . We might associate a random walk $(S_{\Lambda}(k))_{k=0}^{\infty}$ with jump variable X_{Λ} with probabilities $p_{\Lambda}(j) \doteq p_{s,s+i}$.

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Importance sampling scheme

Our importance sampling scheme will be a mixture of two sets of exponentially shifted jump probabilities of the jump variables X_{Λ} .

For any $\theta \in \mathbb{R}^2$, the θ -shifted jump X^{θ}_{Λ} has jump probabilities

 $p^{\theta}_{\Lambda}(i) = e^{\langle \theta, j \rangle - \psi_{\Lambda}(\theta)} p_{\Lambda}(i),$

where $\psi_{\Lambda}(\cdot)$ is the log moment generating function of jump variable X_{Λ} .

This gives us a set of 4 jump (or transition) probability densities.

We have two of such sets, and before we simulate a sample path, we choose randomly a set.



The blue arrow indicates the 'natural' drift.

We show the transition rates; for the discrete-time Markov chain these are normalized to probabilities.



A sample path after change of measure



The idea of the importance sampling scheme: until a certain time $n\tau$ it follows the original transition probabilities.

Sample path large deviations

Define continuous processes $(S^{[n]}(t))_{0 \le t \le T}$, n = 1, 2, ..., by scaling:

$$S^{[n]}(t) = S(nt)/n$$
 for $t = 0, 1/n, 2/n \dots, T$,

and linear interpolation in the other points.

Consider absolute continuous functions $\phi : [0, T] \rightarrow \mathbb{R}^2_+$. Then (Ignatiouk 2005)

$$-\lim_{\epsilon \downarrow 0} \lim_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left(\sup_{0 \le t \le T} \left| S^{[n]}(t) - \phi(t) \right| < \epsilon \right) = \int_0^T \ell_{\Lambda(\phi(t))}(\phi'(t)) \, dt \doteq I(\phi),$$

where $\ell_{\Lambda} : \mathbb{R}^2 \to [0, \infty]$ are so-called locate rate functions (see forthcoming slides).

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The local rate functions

The local rate function $\ell_{\Lambda}(\nu)$ is the convex conjugate of a certain convex function $\psi : \mathbb{R}^2 \to \mathbb{R}$:

$$\ell_{\Lambda}(v) = \sup_{\theta \in \mathbb{R}^2} \left(\langle \theta, v \rangle - \psi(\theta)
ight).$$

It can be determined numerically via the method in (Ignatiouk 2001). The optimizer θ_v is called the optimal shift factor associated with speed v.

When θ_{ν} is used to exponentially shift the jump probabilities of the internal jump variable $X_{\{1,2\}}$, the speed vector ν corresponds with the drift of the shifted jump variable $X_{\{1,2\}}^{\theta_{\nu}}$, thus restricted for being a convex combination of the jumps (-1, 0), (0, -1), (1, 1).

Clearly, the speed vectors in the boundary faces $F_{\{1\}}$ and $F_{\{2\}}$ are restricted to be 0 in the perpendicular direction and between -1 and 1 in the parallel direction. (They do not correspond with drifts!)

Denote by V_{Λ} the set of feasible speed vectors v in face F_{Λ} .

Sample path large deviations (cont'd)

Notice that for the fork-join queue problem

$$\gamma_n(x, y, T) = \mathbb{P}\left(S(nT)/n \in D|S(0)/n = x\right) = \mathbb{P}\left(S^{[n]} \in E\right),$$

where *E* is an appropriate set of absolute continuous paths $\phi: [0,T] \to \mathbb{R}^2_+$ with specifically $\phi(0) = x$ and $\phi(T) \in D$ (the rarity set).

Let $\tilde{E} \subset E$ be the subset of piecewise linear paths of the following form.

 $\phi = \phi_{\tau,v}$: it follows the natural drift until time τ and then it goes straight at constant speed $v = \phi'(t)$ to a point in the rarity set *D*.

One can show that

$$\lim_{n\to\infty}\frac{1}{n}\log\,\mathbb{P}\left(S^{[n]}\in E\right)=\lim_{n\to\infty}\frac{1}{n}\log\,\mathbb{P}\left(S^{[n]}\in \tilde{E}\right)=-I(\tilde{E}),$$

where $I(\tilde{E}) = \inf_{\tau,\nu} I(\phi_{\tau,\nu})$, and $I(\phi_{\tau,\nu}) = (T - \tau)\ell_{\Lambda}(\nu)$ assuming that the second part of the path runs entirely in face F_{Λ} .

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The paths with constant speed

We restrict to starting state x = (0, 0).

Consider paths that stay in 0 during τ time units.

Let $V_{\Lambda}(\tau)$ be the set of feasible speed vectors v in face F_{Λ} such that $(T - \tau)v \in D$, i.e.,

$$V_{\Lambda}(\tau) = \{ v \in V_{\Lambda} : \phi_{\tau,v} \in \tilde{E} \}.$$

Hence, there is a one-to-one correspondence

$$\begin{split} \tilde{E} &\leftrightarrow \bigcup_{\tau \geq 0} V_{\{1,2\}}(\tau) \cup \bigcup_{\tau \geq 0} V_{\{1\}}(\tau) \cup \bigcup_{\tau \geq 0} V_{\{2\}}(\tau) \\ &= V_{\{1,2\}}(0) \cup V_{\{1\}}(0) \cup V_{\{2\}}(0) \doteq V(0). \end{split}$$

Efficient importance sampling

The idea is to choose (better: find) a set of speeds $v^{(1)}, \ldots, v^{(m)}$ and associated optimal shift factors $\theta^{(1)}, \ldots, \theta^{(m)}$, such that

$$V(0) \subset \bigcup_{i=1}^m \mathcal{H}(v^{(i)}),$$

where

 $\mathcal{H}(v) = \{ w \in V(0) : \langle \theta_v, w \rangle \ge \langle \theta_v, v \rangle \}.$

Then, any mixture importance sampling scheme with exponentially shifted probability densities using shift factors $\theta^{(1)}, \ldots, \theta^{(m)}$ is asymptotically optimal (Bucklew 1990, 2004).

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Example (cont'd)

Results for scalings n = 25-500 with sample size k = 50000 for plotting the relative half width of the 95% confidence interval for estimator $\hat{\gamma}_n$,

 $\mathsf{RHW} = 1.96\sqrt{\mathsf{Var}[\widehat{\gamma}_n]}/\mathbb{E}[\widehat{\gamma}_n],$

and ratio $RAT = \log \mathbb{E}[(\widehat{\gamma_n})^2] / \log \mathbb{E}[\widehat{\gamma_n}]$. (efficient estimators have RAT that converge to 2).



Example

 $\lambda = 1, \mu_1 = 1.5, \mu_2 = 2, x = (0, 0), y = (1, 1.2), T = 10.$

We were able to find a solution of two speed vectors such that

$$V(2) \doteq \bigcup_{\tau \ge 2} V_{\{1,2\}}(\tau) \cup \bigcup_{\tau \ge 2} V_{\{1\}}(\tau) \cup \bigcup_{\tau \ge 2} V_{\{2\}}(\tau) \subset \mathcal{H}(\nu^{(1)}) \cup \mathcal{H}(\nu^{(2)}).$$

We mix them 0.8 (red path) and 0.2 (blue path).



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Conclusion and further research

- We have developed an importance sampling scheme which is a mixture of two time-dependent (state-independent) exponentially shifted densities.
- Excellent simulation results.
- Need to prove that using *V*(2) in stead of *V*(0) still gives asymptotical optimality. (The large deviations asymptotic still holds.)
- Further investigations include other starting points, and other algorithms, for instance with mixing transition probabilities that depend on state and time.

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